

How does bank competition affect credit risk? Evidence from loan-level data

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Abstract

This paper studies how loan credit risk depends on competition in the banking sector. We estimate an empirical model of credit risk using data from the Spanish Credit Register on individual loans to non-financial firms in 1992–2007. Our results show that credit risk decreases with the level of competition in the credit market, and they are consistent with the prediction from the moral hazard view on the determinants of credit risk. We also find that the probability of loan default varies with characteristics of the bank, the local market and macro variables.

JEL classification: G21

Keywords: Banks, moral hazard, credit risk, competition, product differentiation, credit register.

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1. Introduction

Whether or not more competition in banking markets increases the risk exposure of banks and, through contagion, contributes to the fragility of the whole banking system is a relevant but unsettled research and policy question (Vives, 2016). There are two main competing views, i.e., the *charter value* and *moral hazard*, on the relation between banking competition, risk exposure of individual banks and financial stability. The *charter value* predicts that more intense competition in the deposits market has a negative effect on financial stability (Keeley, 1990; Hellmann et al., 2000), while the *moral hazard* view predicts that market competition contributes to lower credit risk and more financial stability (Stiglitz and Weiss, 1981; Caminal and Matutes, 2002; Boyd and De Nicoló, 2005). The different theoretical predictions have motivated empirical research to test which one holds in data (Kick and Prieto, 2014; Forssbäck and Shehzad, 2015). However, the different scope of analysis and the different databases complicate the comparison of the results.

This paper presents new empirical evidence on the relationship between bank risk-taking behavior and bank competition, with two important novelties: the use of loan level data and the analysis of the competition effect at local-market level. First, the use of loan level data has clear advantages over bank or country level data, because granting a loan is, *per se*, a risk-taking decision. Second, the structural measure of market competition in our paper is the density of bank branches in spatially delimited markets, where transportation costs generate spatial differentiation that decreases with the density of branches. In doing so, we acknowledge that the relevant markets for retail banking services are local markets served by bank branches.

2. The empirical model and hypotheses

The econometric model to be estimated is formulated as follows:

$$\begin{aligned} \text{Default}_{ijpt} = & \text{Compet}'_{pt}\beta_C + \text{Prov}'_{pt}\beta_P + \text{Bank}'_{jt}\beta_B + \text{Macro}'_t\beta_E + \\ & + \text{Loan}'_{it}\beta_L + \gamma_p + \lambda_j + \varepsilon_{ijpt} \end{aligned} \quad [1]$$

Default_{ijpt} is a dummy that takes the value of 1 if the loan granted to firm i by bank j in province p defaults in year t , and zero otherwise. The vector Compet'_{pt} includes variables that capture the intensity of competition from spatial differentiation in the local market;

$Prov'_{pt}$ includes characteristics of the borrowers in the province p that capture their credit quality. $Bank'_{jt}$ is a vector of variables that refer to the bank j that grants the loan, and $Macro_t$ includes variables of the general economic conditions that are common to all loans at time t . $Loan$ is a vector of dummy variables that controls for the characteristics of the loan (type, maturity, collateral). The γ_p and λ_j are province and bank fixed effects, respectively, and ε_{ijpt} is the random disturbance term. Table 1 show the definition of the main variables used in the analysis.

From the assumption of spatial positive transportation costs, higher density of branches in the province implies less spatial differentiation and more competition. Then, we can test the two competing hypotheses: Under the moral hazard (charter value) view, the probability of loan default decreases (increases) with the density of branches in the province. Since with free opening and closing of branches the density of branches is a function of the size of the market demand, the competing hypotheses are also tested with market demand variables, such as population and per-capita income.

To compare the results with those in the previous literature and highlight the relevance of the market definition and the use of loan level data, we add the concentration of banks in the market (HHI calculated with the number of branches in a province) as explanatory variable of the probability of loan default. Now, more concentration implies less competition, so the expected sign of the coefficient of HHI is the opposite of the sign of the density of branches.

The vector $Macro_t$ includes the interbank interest rate as a relevant variable for testing the competing hypotheses. Higher market interest rate increases the loan interest rate for a given level of competition, and the moral hazard view predicts that this would increase the probability of loan default (monetary transmission). The charter value view has no direct prediction on the effect of the market interest rate, but there is an indirect effect through the deposit market. Since higher interbank interest rates increase the profits in the deposit market, the value of the charter value of the bank increases. As a result, the bank becomes more conservative and the risk-appetite of banks decreases, reducing the probability of loan default.

3. Database

The data on bank loans to business firms comes from the Spanish Credit Register (CIR). This unique database contains information on all bank loans granted to nonfinancial firms in Spain above the threshold of 6,000 euros. For each bank loan, we know the bank that

granted it; the province where the borrower resides; and some loan characteristics such as the type of loan, the maturity, the guarantees, and if the loan is in default or not. We restrict our sample to loans lower than 1 million euros to focus on firms borrowing from branches in local markets, that is, 95% of the population of business loans in the CIR database. Larger loans are granted to large firms whose relevant market is likely to be the credit office of the parent bank, rather than a local market around a particular branch.

The relevant market for the retail activities of a bank is the province where it grants the loan. The data on province level variables comes mainly from the Spanish Institute of Statistics (INE). The number of branches and bank level variables come from the accounting statements that each bank reports regularly to the Banco de España. Descriptive statistics of the values of the variables of the empirical model are presented in columns three and on in Table 1.

4. Empirical results

The results of the estimation of different specifications of model (1) appear in Table 2. All the models are estimated with standard errors robust to heteroscedasticity and clustered at bank-province level.

Column I present the results of the estimation with the number of bank branches, $\ln NBRANCH$, as the only market competition variable. The estimated coefficient is negative and statistically significant, consistent with the moral hazard view. The positive and significant coefficient of $INTERBANK$ is also consistent with the prediction from the moral hazard view that higher interest rates induce higher credit risk.

The probability of loan default is also lower in provinces with higher GDP growth rate and lower $UNEMPLOYMENT$, that is, lower credit risk in more prosperous provinces. Loans granted by less productive banks (lower $PROD$) and with looser credit policies (high credit growth $LOANGR$ and NPL) are riskier than otherwise. Finally, the capitalization of banks has a negative effect in the probability of loan default (negative coefficient of $CAPITAL$) indicating no evidence of substitution between financial and credit risk.

Column II includes market concentration HHI and its square, $HHISQ$, as explanatory variables, and time dummies replace macro variables to strengthen the control for time varying effects. The estimated coefficients of HHI and $HHISQ$ are not statistically significant, and the coefficient of $\ln NBRANCH$ increases its magnitude and statistical significance, compared to Column I. Thus, the market concentration variable strengthens, not diminishes, the competition effect of the density of branches. Controlling

for the distribution of the number of branches across banks, an increase in the density of branches implies more branches per bank. Thus, the results indicate that the competition effect in credit risk from increasing the density of bank branches is higher when the increase is evenly distributed among banks. The results also consistent with an increase in both the number of banks and the number of branches. In fact, the correlation of the number of branches and the number of banks in the province is relatively high in the data (88%).

Column III of Table 2 shows the results of the estimation of the model when the density of bank branches is replaced by the proxy variables of market size, *GDPCAP* and *POPULATION* (both in logs). The estimated coefficients of the two variables are negative as expected if the size of the market is a proxy of intensity of competition because they will attract the entry of larger number of branches.

5. Conclusion

Credit risk is at the core of the risk exposure of individual banks and, through contagion, also at the core of financial stability. The way in which competition in loan and deposit markets affects the risks taken in the lending decisions of banks is particularly important. If franchise value theory holds and more market competition induces riskier behavior by banks, then pro-competition policies and prudential banking regulation might come into conflict. In contrast, if the moral hazard view holds, market competition is aligned with safer bank loans.

The empirical results with data on individual loans granted to business firms by Spanish banks broadly support the moral hazard view that more competition decreases credit risk. More concretely, we find a negative effect of the number of branches on the probability of default, which might also be explained by the higher number of banks operating in provinces with a higher number of branches (correlation of 88%). Bank market concentration does not directly affect credit risk, though the negative effect of the density of branches on credit risk is higher if branches are more evenly distributed among banks. The positive effect of the interbank interest rate in the probability of default also supports the moral hazard hypothesis.

The empirical results also show that banks with a higher equity ratio grant loans with a lower probability of default. There is no evidence of a trade-off between lower financial risk and higher credit risk. Rather, the empirical evidence is consistent with the prediction from the charter value that the option value of deposit insurance decreases with

the capital ratio and for this reason more capitalized banks have lower incentives to take higher credit risks).

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Table 1. Descriptive statistics of the main variables

		MEAN	SD	P10 th	P50 th	P90 th
DEPENDENT VARIABLE						
<i>DEFAULT</i>	=1 if the loan is defaulted	0.03	0.18	0.00	0.00	0.00
INDEPENDENT VARIABLES						
<i>Competition Variables</i>						
<i>lnNBRANCH</i>	Log of the number of branches in the province	6.09	0.93	5.18	6.10	6.95
<i>lnPOPU</i>	Log of the total population	16.17	0.86	15.21	16.17	17.16
<i>GDPCAP_p</i> (th€)	Gross Domestic Product per capita, constant Euros of 2000	17.96	6.69	9.94	17.11	27.18
<i>Province Variables</i>						
<i>GDPG_p</i> (%)	GDP growth at constant euros of the province	0.03	0.02	0.00	0.03	0.06
<i>UNEMP_p</i> (%)	Unemployment rate of the province	0.07	0.02	0.04	0.07	0.11
<i>NPL_p</i>	Ratio of non-performing loans in the province	0.05	0.06	0.00	0.05	0.13
<i>RURAL</i> (%)	Percentage of population in villages smaller than 2000 inhab	14.95	14.34	0.67	11.13	37.01
<i>Bank Variables</i>						
<i>PROD</i>	Productivity of the bank from Levinsohn and Petrin methodology as in Martín-Oliver et al. (2013)	7.58	0.45	7.03	7.58	8.19
<i>lnASSETS</i>	Log of the total assets of the bank	15.55	1.54	13.60	15.55	17.69
<i>LOANGR</i>	Yearly growth rate of the volume of loans of the bank	0.16	0.11	0.00	0.16	0.31
<i>NPL_j</i>	Ratio of Non-performing loans	0.05	0.10	0.00	0.05	0.13
<i>CAPITAL</i>	Sum of equity and reserves over total assets of the bank	0.05	0.03	0.03	0.05	0.09
<i>Time-varying variables</i>						
<i>INTERBANK</i>	Euribor 12 months, nominal terms	0.06	0.03	0.02	0.06	0.11
<i>GDPG</i>	Spanish GDP growth at constant euros	0.03	0.01	0.01	0.03	0.05
<i>INFLATION</i>	Growth rate of Consumption Price Index	0.03	0.01	0.02	0.03	0.05
<i>Concentration</i>						
<i>HHI</i>	Herfindahl-Hirschman index for the branch network in a province	0.15	0.05	0.10	0.15	0.21

Table 2. Determinants of Credit Risk in Bank Loans to Business Firms

Dep =1 if Loan is Default	II	II	III
<i>Competition Variables</i>			
$\ln NBRANCH_{t-1}$	-0.017 ** (0.008)	-0.026 *** (0.010)	
$\ln POPU_{t-1}$			-0.025 * (0.014)
$\ln GDPCAP_{p,t-1}$			-0.036 ** (0.014)
<i>Province Variables</i>			
$GDPG_{p,t-1}$	-0.043 *** (0.016)	-0.037 ** (0.015)	-0.023 (0.017)
$UNEMP_{p,t-1}$	0.096 ** (0.047)	0.144 ** (0.067)	0.119 * (0.070)
$NPL_{p,t-1}$	0.030 (0.031)	0.030 (0.035)	0.024 (0.034)
$RURAL_{p,t-1}$	0.013 (0.025)	0.016 (0.025)	0.061 * (0.034)
<i>Bank Variables</i>			
$PROD_{t-1}$	-0.010 ** (0.004)	-0.010 *** (0.004)	-0.009 ** (0.004)
$\ln ASSETS_{t-1}$	0.003 (0.003)	0.003 (0.004)	0.000 (0.004)
$LOANGR_{t-2}$	0.009 *** (0.003)	0.005 * (0.003)	0.005 * (0.003)
$LOANGR_{t-3}$	0.016 *** (0.003)	0.015 *** (0.003)	0.016 *** (0.003)
$NPL_{j,t-1}$	0.424 *** (0.050)	0.425 *** (0.052)	0.429 *** (0.049)
$CAPITAL_{t-1}$	-0.115 ** (0.051)	-0.112 (0.051) **	-0.138 *** (0.048)
<i>Time-varying variables</i>			
$INTERBANK_{t-1}$	0.306 *** (0.054)		
$GDPG_{t-1}$	0.036 (0.036)		
$INFLATION_{t-1}$	0.081 (0.072)		
<i>Concentration</i>			
HHI_{pt-1}		0.011 (0.148)	
$HHISQ_{pt-1}$		-0.396 (0.389)	
Time Dummies	NO	YES	YES
Fixed Effects	Province, Bank, Type Loan, Type Collateral, Maturity	Province, Bank, Type Loan, Type Collateral, Maturity	Province, Bank, Type Loan, Type Collateral, Maturity
N.Observations	3,681,057	3,681,057	3,681,057

The *, **, and *** represent significance at the 10%, 5%, and 1% levels respectively. The *t*-ratios are in parentheses. Note: The dependent variable is a dummy that takes the value of one if the loan defaults and zero otherwise. The models are estimated with linear probability models that control for fixed effects, and the standard errors are robust to heteroscedasticity and are clustered at the bank-province level. Definitions for the variables are in Table 1