

Contents lists available at ScienceDirect

Computers in Human Behavior

journal homepage: www.elsevier.com/locate/comphumbeh

Sex, age and cyber-victimization: A meta-analysis.



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ARTICLE INFO

Keywords: Cyberbullying Cyber-victimization Meta-analysis Sex Age Culture

ABSTRACT

Cyberbulling is one of the biggest challenges the school faces. However, the lack of coherence between the data of the literature review makes it necessary to consider which elements are the ones that truly lead to the appearance of cyber-victimization. Through the meta-analysis methodology, it has been tried to clarify the role of sex (k = 41 samples, n = 176,658 adolescents) and age (k = 45 samples, n = 238,977 adolescents) in cyber-victimization. The effect size for the random model is small for both sex (r = 0.058; p < 0.00, 95% CI = 0.090; 3.45) and for age (r = 0.094; p = 0.004; 95% CI = 0.015; 2.910). Indications of significant differences in sex are observed, with women being the most affected. However, the results of the meta-regression have shown how the North American culture plays a key role in age as a moderating variable in relation to the rest of cother-victimization. More specifically, there is a positive relationship between age and cybervictimization, so that the older the age, the higher the cybervictimization, but this is negatively mediated by the American culture. At the same time, some socio-contextual characteristics also seem to have effects on this aspect. Considering this, some important practical implications emerge related to the need to address the study, care and prevention of cyber-victimization as well as any form of violence that occurs inside and outside the classroom.

1. Introduction

Cyberbullying is a type of violence between equals carried out using new information and communication technologies. These tools, in turn, allow aggressive behaviour to be maintained at any time (Garaigordobil, 2011). However, data on prevalence are inclusive, as it varies greatly according to the instrument used and nationality (Romera et al., 2016; Zych et al., 2016). Variability in diagnosis is primarily mediated by cultural factors related to the macro system (Bronfenbrenner, 1979), i.e. cultural norms, social responses and protective values. The macrosystem is fundamental when it comes to inhibiting or favouring cyberbullying (Baldry et al., 2015), as there are important differences in the actions carried out by different cultures (Lozano-Blasco et al., 2020).

In addition, there are a number of social factors, both external and internal, which mediate its development. In the first place, the irruption into daily life of social networks increases the probability of both suffering cyber-victimization and playing the role of cyber-stalker, mediated by the number of hours spent on the Internet (Choi et al., 2019; Marciano et al., 2020; Shapka et al., 2018; Tsitsika et al., 2015), since victims of this type of violence present very high scores CIU (Compulsive Internet Use) and TOB (Troubled Offline Behaviour) (Wachs et al., 2018). On the other hand, the family has a key role in the development of cyberbullying. Factors such as divorce, the mother's low educational level, the father's unemployment (Charalampous et al., 2018; Chen et al., 2018) or "incivility" (Bai et al., 2020) correlate with cyber-victimization. On the other hand, family control is essential, i.e. clear and precise rules on acceptable or unacceptable behaviour. However, there are also cultural differences in this area, with Polish families being the most successful in this respect (Athanasiou et al., 2018). An authoritarian and inconsistent parental style in controlling the Internet is related to being a cyber-victim, and cyber-aggressor (Katz et al., 2019). Furthermore, cyberbullying correlates with internalising problems in such a way that suffering from cyber-victimization alters psychological health and decreases self-esteem and perceived effectiveness (Alvarez-Garcia et al., 2015; Holfeld & Mishna, 2019; Lei et al., 2019; Sampasa-Kanyinga et al., 2020; Tsitsika et al., 2015; Waasdorp & Bradshaw, 2015; Wolke et al., 2017). Similarly, some studies indicate that victims of cyberbullying have difficulty expressing their emotions

https://doi.org/10.1016/j.chb.2022.107491

Received 10 February 2022; Received in revised form 6 September 2022; Accepted 17 September 2022 Available online 28 September 2022

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(Extremera et al., 2018) and, more specifically, with anger (Lonigro et al., 2015).

With regard to the sociodemographic variables of sex and age, there is no agreement in the scientific community about their role in cybervictimization. Sex is an inconsistent variable since, on the one hand, numerous studies point to the non-existence of significant differences (Simckes et al., 2017; Hemphill & Heerde, 2014; Holfeld & Mishna, 2019), while others point to differences that mostly affect women (Álvarez-García et al., 2015; Bauman et al., 2013; Buelga et al., 2017; Cuadrado-Gordillo & Fernández-Antelo, 2016; DeSmet et al., 2014; Extremera et al., 2018; González-Cabrera et al., 2017, 2018; Messias et al., 2014; Moreno-Ruiz et al., 2019). In terms of age, some authors find significant differences and others do not (Athanasiou et al., 2018; Holfeld & Mishna, 2019; Katz et al., 2019; Selkie et al., 2016; Shapka et al., 2018). Among those who advocate the role of age as a mediator are three different approaches:

- a) Cyberbullying decreases with increasing age (Jang et al., 2014; Lonigro et al., 2015; Waasdorp & Bradshaw, 2015),
- b) Cyberbullying increases with increasing age (Bauman et al., 2013; Chen et al., 2018; Choi et al., 2019; Festl & Quandt, 2016; Messias et al., 2014; Simckes et al., 2017) and
- c) Cyberbullying experiences a curvilinear development that peaks in mid-adolescence (14–16 years) (Gámez-Guadix et al., 2015; Sakellariou et al., 2012; Tokunaga, 2010; Wolke et al., 2017).

2. Methodology

2.1. Research question

The role of sex and age in cyber-victimization presents disparate data in the scientific literature. The present meta-analyses investigate the significance of sex and age differences in cyber-victimization and the possible existence of moderators.

2.2. Inclusion and exclusion criteria

Regarding the selection of studies, this meta-analysis was carried out following the manual of systematic reviews of Cochrane in Higgins and Green (2011). The inclusion criteria were:

- A. a) the age of the participants in each study. Specifically, the average age range of the sample is between 11.5 and 18.9 years to cater for the diversity of academic situations by Izquierdo-Martínez (2007).
- (b) the methodological approach to empirical studies, as they were to be experimental, based on a quantitative method of statistical analysis of data.
- (c) the date of publication between 2013 and 2019 b y Borenstein et al. (2021), Cívico Ariza et al., (2021) and Gurevitch et al. (2018).
- d) intra-observer reliability, those scientific publications with the greatest impact on the subject (Q1 Scimago Journal & Country Rank).

The exclusion criteria following the indications of Botella and Sánchez (2015) and Moreau and Gamble (2020) were:

- a) the presence of Special Educational Needs (SEN) as a main feature of the sample. Even with this, it was determined that those investigations in which the group with SEN were in agreement with the normal curve would be admitted.
- (b) the lack of clarity, precision and rigour in the description of the method of quantitative data analysis adopted. We follow the indications of Hunter and Schmidt (2004) and Friese and Frankenbach (2020)

(c) the absence of a comparative analysis between the two sexes with regard to cyber-victimization or the consideration of only one of them.

2.3. Search strategies

The search strategy followed the parameters of Botella and Sánchez (2015). The search was carried out in October 2019. The Boolean action was: "Cyberbullying AND cyber-victimization OR cybervictim". It was carried out using four databases, establishing the following filters:

- Scopus: "article title, abstract, keywords" and "article".
- Psycinfo: "adolescent population", "academic publications", "exclude dissertations" and "keywords".
- Science Direct: "abstract o key words " y "research articles".
- Pubmed: "title/abstract".

In addition, the temporary range of 2013–2019 was specified. Each article was selected manually, according to the established inclusion and exclusion criteria, as well as reviewing the bibliography contained in them and adding seven researches for the gender variable and six for age (see Fig. 1). It should be noted that these investigations were divided into two blocks: those that dealt with differences in sex and those that dealt with the age variable, generating two meta-analyses.

2.4. Codification process

The registration protocol for this meta-analysis was carried out following the Cochrane systematic review manual in Higgins and Green (2011) and PRISMA-Statement website (). The coding of the studies was developed manually according to the inclusion and exclusion criteria explained above. In the execution of this meta-analysis, the statistical software CMA was used, which allowed the conversion of the values to Fisher Z and at the same time to carry out the publication bias tests (Egger and Begg), to calculate statistics on heterogeneity, meta-regressions and model comparisons, as well as to obtain figures such as the Forest Plot, funnel and Fisher Z meta-regression graphs.

3. Results

3.1. Description of the sample

Before going on to describe the results in terms of gender and age differences in the rate of cyber-victimization in recent years (2013–2019), it should be noted that four studies do not report data on the rate of men and women and, similarly, some do not present the average age, but a range of years. In this case, the arithmetic average was used to calculate it, resulting in 14.64 years.

In total, the meta-analysis on gender differences in the rate of cybervictimization is made up of a total sample of 176,658 participants (k = 41) from 32 studies (Table 1). In terms of sex, 64,260 are men and 68,417 are women. The study with the least number of participants had 60 and the largest had 28,104. The European population (Spain, Italy, Cyprus, Germany, Belgium, Israel, Portugal and studies with samples from different European nations) represents 58.88%. The American sample (USA and Canada) covered 12.99% of the total. Asian culture (China, South Korea) accounted for 26.29%, while African culture (Tanzania and Nigeria) with 0.69% represented a small percentage. As in the previous case, 1.15% belonged to a sample made up of various transcontinental nations.

According to the age differences, a total sample of 238,977 individuals (k = 45) from 32 studies was obtained (Table 1). In terms of gender, 85,012 are men and 89,097 are women. The smallest sample was 175 and the largest 28,104. Within the cultural diversity of the sample, Europe (Spain, Portugal, UK, Italy, Israel, Belgium, Germany, Cyprus and others with samples from different European nations)

| | Cyberbullying. and Cyber- | SEARCHING victimisation or cyber-victim - 2019 | |
|---|---|---|--|
| SCOPUS * Article title, abstract, keywords. * Article | PUBMED * Title / Abstract | PSYCO INFO * Young adults * Research articles * Dissertations exclusion * Keywords | SCIENCE DIRECT * Abstract or keywords. * Research articles |
| AGE VARIABLE | AGE VARIABLE | AGE VARIABLE | AGE VARIABLE |
| TOTAL: 109 <i>Exclusion criteria:</i> Average age: 4 Methods: 31 Theme: 4 Rigorous methodology: 26 Inaccessibility: 25 Statistical absences: 3 | TOTAL: 44 Exclusion criteria: Average age: 2 Methods: 9 Rigorous methodology: 9 Inaccessibility: 4 Repeat: 10 Extreme values: 1 | TOTAL: 29 Exclusion criteria: Methods: 8 Statistical absences: 12 Inaccessibility: 5 | TOTAL: 57 <i>Exclusion criteria:</i> Methods: 4 Rigorous methodology: 5 Repeat: 7 Extreme values: 2 Statistical absences: 2 No data: 29 |
| SELECTED: 12 | SELECTED: 2 | SELECTED: 3 | SELECTED: 8 |
| | Added by bibliog | raphic references: 6 | |
| SEX VARIABLE | SEX VARIABLE | SEX VARIABLE | SEX VARIABLE |
| TOTAL: 109 Exclusion criteria: Average age: 8 Methods: 33 Extreme values: 1 Rigorous methodology: 26 Inaccessibility: 25 Statistical absences: 6 | TOTAL: 44 <i>Exclusion criteria:</i> Average age: 2 Methods: 12 Rigorous methodology: 9 Inaccessibility: 4 Repeat: 10 CMA Mistake: 1 | TOTAL: 29 Exclusion criteria: Methods: 8 Statistical absences: 10 Inaccessibility: 4 | TOTAL: 57 <i>Exclusion criteria:</i> Methods: 2 Rigorous methodology: 5 Repeat: 7 Extreme values: 2 Statistical absences: 29 |
| SELECTED: 9 | SELECTED: 2 | SELECTED: 7 | SELECTED: 10 |
| | Added by bibliog | raphic references: 7 | |

Fig. 1. Flow chart of the study search and selection process.

Table 1

Socio-demographic description of both variables.

| SEX VARIABLE | | | | | |
|---|---|---|--|---|---|
| Study | Average age | Sample | Men | Women | Nation |
| Athanasiou et al. (2018) (a,b,c y d) | 15,5 | 13,708 | 5786 | 6586 | Europea |
| Baldry, Farrington, & Sorrentino, 2017 | 15,4 | 2785 | 1320 | 1465 | Italy |
| Baldry et al. (2019) | 15,5 | 4390 | 2016 | 2374 | Italy |
| Bauman et al. (2013) (a y b) | 16,5 | 1482 | 757 | 725 | US |
| Baumgartner et al. (2014) | 13,5 | 14,946 | 7428 | 7518 | Europe |
| Buelga et al. (2017) | 14,5 | 1062 | 546 | 516 | Spain |
| | | | | | - |
| Chang et al. (2013) | 15 | 2992 | 1555 | 1437 | China |
| Charalampous et al. (2018) | 11,72 | 868 | 410 | 451 | Cyprus |
| Chen et al. (2018) (b) | 15,86 | 18,341 | 9776 | 8565 | China |
| Choi et al. (2019) | 15 | 7109 | 3349 | 3760 | South Korea |
| Cuadrado and Fernández 2016 | 14,1 | 1648 | 843 | 805 | Spain |
| DeSmet et al. (2014) | 15,3 | 204 | 78 | 126 | Belgium |
| DeSmet et al. (2018) | 15 | 1307 | _ | _ | Belgium |
| Erreygers et al. (2016) | 12,6 | 2309 | 1161 | 1148 | Belgium |
| Extremera et al. (2018) | 14,1 | 1660 | 824 | 836 | Spain |
| | | | | | - |
| Ferrer-Cascales et al. (2019) | 13,08 | 2057 | 1036 | 1021 | Spain |
| Gámez-Guadix et al. (2015) | 14,8 | 680 | 270 | 410 | Spain |
| Giménez Gualdo et al., (2015) | 14,77 | 4353 | 639 | 714 | Spain |
| González-Cabrera et al. (2017) (a) | 14,89 | 371 | 191 | 180 | Spain |
| González-Cabrera et al. (2017) (b) | 15,58 | 60 | 26 | 34 | Spain |
| Holfeld and Mishna (2019) (a) | 13,7 | 510 | 196 | 314 | Canada |
| Holfeld and Mishna (2019) (b) | 14,63 | 422 | 162 | 260 | Canada |
| Holfeld and Mishna (2019) (c) | 15,76 | 329 | 133 | 196 | Canada |
| Jang et al. (2014) | 14 | 16,190 | - | - | South Korea |
| Jetelina et al. (2019) | | 4297 | | | US |
| | 15,5 | | - | - | |
| Katz et al. (2019) | 13,25 | 175 | 86 | 89 | Israel |
| Messias et al. (2014) | 16 | 13,846 | - | - | US |
| Moreno-Ruiz et al. (2019) | 14,69 | 2399 | 1024 | 1375 | Spain |
| Olumide et al. (2016) | 14,2 | 653 | 318 | 335 | Nigeria |
| Quintana et al. (2018) | 14,1 | 1650 | 825 | 840 | Spain |
| Shapka et al. (2018) (a) | 14,25 | 426 | 239 | 187 | Tanzania |
| Shapka et al. (2018) (b) | 12,64 | 594 | 292 | 302 | Canada |
| Veiga Simão et al., (2017) | 13,6 | 3525 | 1683 | 1837 | Portugal |
| Wachs et al. (2018) | | 2042 | 946 | 1096 | Dutch, German, Thai, and U |
| | 14,2 | | | | |
| Wolke et al. (2017) | 13,5 | 2745 | 1184 | 1561 | UK |
| Wong et al. (2014) | 13,36 | 1917 | 1046 | 871 | China |
| AGE VARIABLE | | | | | |
| Álvarez-García et al., (2015) (a y b) | 15 | 3180 | 1542 | 1632 | Spain |
| Athanasiou et al. (2018) (a,b,c,d) | 15,5 | 13,708 | 5786 | 6586 | Europea |
| *Barboza (2015) | 14,77 | 5589 | 2850 | 2739 | US |
| Bauman et al. (2013) (a y b) | 16,5 | 1482 | 757 | 725 | US |
| Buelga et al. (2017) | | 1062 | 546 | 516 | Spain |
| | 14,5 | | | | - |
| Charalampous et al. (2018) | 11,72 | 868 | 410 | 451 | Cyprus |
| Chen et al. (2018) (b) | 15,86 | 18,341 | 9776 | 8565 | China |
| Choi et al. (2019) | 15 | 7109 | 3349 | 3760 | South Korea |
| DeSmet et al. (2014) | 15,3 | 204 | 78 | 126 | Belgium |
| DeSmet et al. (2018) | 15 | 1307 | - | - | Belgium |
| Erreygers et al. (2016) | 12,6 | 2309 | 1161 | 1148 | Belgium |
| Extremera et al. (2018) | 14,1 | 1660 | 824 | 836 | Spain |
| Ferrer-Cascales et al. (2019) | 13,08 | 2057 | 1036 | 1021 | Spain |
| Festl and Quandt (2016) | 15,00 | 1817 | 800 | 1021 | Germany |
| | 15 | | | | - |
| Gámez-Guadix et al. (2015) | | | | 410 | Spain |
| | 14,8 | 680 | 270 | | |
| | 14,77 | 4353 | 639 | 714 | Spain |
| | | | | 714 180 | Spain Spain |
| González-Cabrera et al. (2017) | 14,77 | 4353 | 639 | | - |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) | 14,77 14,89 | 4353 371 | 639 191 | 180 | Spain |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) | 14,77 14,89 13,36 16 | 4353 371 920 658 | 639 191 450 299 | 180 470 363 | Spain Spain Australia |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) | 14,77 14,89 13,36 16 13,7 | 4353 371 920 658 510 | 639 191 450 299 196 | 180 470 363 314 | Spain Spain Australia Canada |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) | 14,77 14,89 13,36 16 13,7 14,63 | 4353 371 920 658 510 422 | 639 191 450 299 196 162 | 180 470 363 314 260 | Spain Spain Australia Canada Canada |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) | 14,77 14,89 13,36 16 13,7 14,63 15,76 | 4353 371 920 658 510 422 329 | 639 191 450 299 196 162 133 | 180 470 363 314 260 196 | Spain Spain Australia Canada Canada Canada |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 | 4353 371 920 658 510 422 329 16,190 | 639 191 450 299 196 162 133 | 180 470 363 314 260 196 | Spain Spain Australia Canada Canada Canada South Korea |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 | 4353 371 920 658 510 422 329 16,190 4297 | 639 191 450 299 196 162 133 - | 180 470 363 314 260 196 - - | Spain Spain Australia Canada Canada Canada South Korea US |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 13,25 | 4353 371 920 658 510 422 329 16,190 4297 175 | 639 191 450 299 196 162 133 - - 86 | 180 470 363 314 260 196 - - 89 | Spain Spain Australia Canada Canada South Korea US Israel |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 | 4353 371 920 658 510 422 329 16,190 4297 | 639 191 450 299 196 162 133 - | 180 470 363 314 260 196 - - | Spain Spain Australia Canada Canada Canada South Korea US |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) Lonigro et al., (2014) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 13,25 | 4353 371 920 658 510 422 329 16,190 4297 175 | 639 191 450 299 196 162 133 - - 86 | 180 470 363 314 260 196 - - 89 | Spain Spain Australia Canada Canada South Korea US Israel |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) Lonigro et al., (2014) Messias et al. (2014) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 13,25 14,91 16 | 4353 371 920 658 510 422 329 16,190 4297 175 716 13,846 | 639 191 450 299 196 162 133 - - 86 324 | 180 470 363 314 260 196 - - 89 392 | Spain Spain Australia Canada Canada Canada South Korea US Israel Italy US |
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| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) Lonigro et al., (2014) Messias et al. (2014) Olumide et al. (2016) Quintana-Orts and Rey (2018) Shapka et al. (2018) (b) *Simckes et al. (2017) (a y b) | 14,77 14,89 13,36 16 13,7 14,63 15,76 14 15,5 13,25 14,91 16 14,2 14,1 14,25 12,64 | 4353 371 920 658 510 422 329 16,190 4297 175 716 13,846 653 1650 426 594 | 639 191 450 299 196 162 133 - - 86 324 - 318 825 239 292 | 180 470 363 314 260 196 - - 89 392 - 335 840 187 302 | Spain Spain Australia Canada Canada South Korea US Israel Italy US Nigeria Spain Tanzania Canada |
| González-Cabrera et al. (2017) González-Cabrera et al. (2018) (2 samples) Hemphill and Heerde (2014) Holfeld and Mishna (2019) (a) Holfeld and Mishna (2019) (b) Holfeld and Mishna (2019) (c) Jang et al. (2014) Jetelina et al. (2019) Katz et al. (2019) Lonigro et al., (2014) Messias et al. (2014) Olumide et al. (2016) Quintana-Orts and Rey (2018) | 14,77 $14,89$ $13,36$ 16 $13,7$ $14,63$ $15,76$ 14 $15,5$ $13,25$ $14,91$ 16 $14,2$ $14,1$ $14,25$ $12,64$ 16 | 4353 371 920 658 510 422 329 16,190 4297 175 716 13,846 653 1650 426 594 10,704 | 639 191 450 299 196 162 133 - - 86 324 - 318 825 239 292 - | 180 470 363 314 260 196 - - 89 392 - 335 840 187 302 - | Spain Spain Australia Canada Canada South Korea US Israel Italy US Nigeria Spain Tanzania Canada US |

(continued on next page)

Table 1 (continued)

| fuble f (continued) | | | | | |
|--------------------------------------|------|------|------|------|-------------------------------|
| Wachs et al. (2018) | 14,2 | 2042 | 946 | 1096 | Dutch, German, Thai, and U.S. |
| Williford et al. (2013) (a,b,c, y d) | 14 | 9914 | 4868 | 5046 | Finland |
| Wolke et al. (2017) | 13,5 | 2745 | 1184 | 1561 | UK |

represents 48.46%. The American sample (USA and Canada) is 32.57%. The Asian sample (China, South Korea) is 17.61%, Africa (Tanzania) is 0.24%, the oceanic sample is 0.27% and the international sample (study of several nations) is 0.85%.

3.2. Size of the effect and statistical significance

The size of the effect was calculated from correlation coefficients, Odds Ratio, Xi2, t, etc. All these were transformed to Fisher's Z values, to ensure that the variance of the effect size was based on the sample size (Martin-Andrés & Luna del Castillo, 2004). Despite the methodological and statistical rigour, two studies were eliminated in the data processing phase with regard to the age variable: Barboza (2015) and Simckes et al. However, they were taken into account for the discussion. Thus, the effect size returned was small, although significant (Cohen, 2013), both for sex differences (r = 0.058; p < 0.00, 95% CI = 0.090; 3.45) and for age differences (r = 0.094; p = 0.004; 95% CI = 0.015; 2.910) (Figs. 2 and 3).

3.3. Statistical analysis

The heterogeneity of the sample was studied as indicated by Cochrane in Higgins and Green (2011). DerSimonian and Laird's (1986) Q statistic for sex (Q = 824,850, gl = 40, p < 0.001) and for age (Q = 592,775, gl = 44, p < 0.001) showed a high variability, rejecting the homogeneity hypothesis. On the other hand, the I2 statistic returned a high value of heterogeneity (Higgins et al., 2003), both for sex (95.15%) and age (92.58%). In other words, variability is caused by heterogeneity and not by chance. Due to this factor, the choice was made to follow the Random Effects model (Martin-Andrés & Luna del Castillo, 2004).

The need to ensure the existence or not of publication bias was stipulated (Botella & Sánchez, 2015; Higgins & Green, 2011;

Meta Analysis

| Study name | | Statistic | s for each | study | |
|----------------------------------|-------------|----------------|----------------|---------|---------|
| | Correlation | Lower limit | Upper limit | Z-Value | p-Value |
| Athanasiou et al.a (2018) | -0,045 | -0,108 | 0,018 | -1,390 | 0,165 |
| Athanasiou et al.b (2018) | -0,118 | -0,207 | -0,027 | -2,528 | 0,011 |
| Athanasiou et al.c (2018) | -0,061 | -0,121 | -0,001 | -2,008 | 0,045 |
| Athanasiou et al.d (2018) | 0,029 | -0,040 | 0,097 | 0,823 | 0,410 |
| Baldry et al. (2019) | -0,029 | -0,064 | 0,006 | -1,617 | 0,106 |
| Baldry et al.b (2017) | 0,171 | 0,124 | 0,218 | 7,032 | 0,000 |
| Bauman et al. b (2013) | -0,051 | -0,108 | 0,006 | -1,761 | 0,078 |
| Bauman et al. a (2013) | 0,087 | 0,036 | 0,137 | 3,353 | 0,001 |
| Baumgartner et al. (2014) | -0,050 | -0,109 | 0,010 | -1,638 | 0,101 |
| Buelga et al. (2017) | 0,092 | 0,032 | 0,151 | 3,006 | 0,003 |
| Chang et al. (2013) | 0,123 | 0,038 | 0,207 | 2,819 | 0,005 |
| Charalampous et al. (2018) | 0,120 | 0,054 | 0,185 | 3,546 | 0,000 |
| Chen et al. b (2018) | -0,026 | -0,056 | 0,004 | -1,710 | 0,087 |
| Choi et al. (2019) | 0,076 | 0,053 | 0,099 | 6,419 | 0,000 |
| Cuadrado and Femández (2016) | 0,502 | 0,465 | 0,537 | 22,387 | 0,000 |
| DeSmet et al. (2014) | 0,045 | -0,203 | 0,287 | 0,350 | 0,726 |
| DeSmet et al. (2018) | -0,094 | -0,222 | 0,038 | -1,402 | 0,161 |
| Erreygers et al. (2016) | 0,060 | 0,019 | 0,101 | 2,885 | 0,004 |
| Ferrer-Cascales et al. (2019) | 0,023 | -0,020 | 0,066 | 1,050 | 0,294 |
| Gámez-Guadix et al. (2015) | 0,087 | 0,012 | 0,161 | 2,277 | 0,023 |
| Giménez et al. (2015) | 0,057 | 0,027 | 0,086 | 3,738 | 0,000 |
| González-Cabrera et al. a (2018) | 0,046 | -0,056 | 0,147 | 0,889 | 0,374 |
| González-Cabrera et al. b (2018) | 0,046 | -0,056 | 0,147 | 0,889 | 0,374 |
| Holfeld & Mishna a (2019) | 0,070 | -0,017 | 0,156 | 1,579 | 0,114 |
| Holfeld & Mishna b (2019) | 0,020 | -0,076 | 0,115 | 0,409 | 0,682 |
| Holfeld & Mishna c (2019) | 0,040 | -0,068 | 0,147 | 0,723 | 0,470 |
| Jang et al. (2014) | 0,143 | 0,105 | 0,181 | 7,228 | 0,000 |
| Jetelina et al. (2019) | 0,055 | 0,027 | 0,083 | 3,810 | 0,000 |
| Katz et al. (2019) | 0,018 | -0,131 | 0,166 | 0,237 | 0,812 |
| Messias et al. (2014) | 0,055 | 0,022 | 0,088 | 3,291 | 0,001 |
| Moreno-Ruiz et al. (2019) | 0,067 | 0,027 | 0,107 | 3,291 | 0,001 |
| Olumide et al. (2019) | 0,127 | 0,031 | 0,220 | 2,603 | 0,009 |
| Pettalia et al | 0,205 | 0,085 | 0,319 | 3,314 | 0,001 |
| Quintana et al.(2018) | -0,067 | -0,115 | -0,019 | -2,737 | 0,006 |
| Shapka et al. a (2018) | 0,070 | -0,025 | 0,164 | 1,442 | 0,149 |
| Shapka et al. b (2018) | 0,010 | -0,071 | 0,090 | 0,243 | 0,808 |
| Veiga et al. | 0,080 | 0,047 | 0,113 | 4,758 | 0,000 |
| Wachs et al. (2018) | 0,086 | 0,043 | 0,129 | 3,893 | 0,000 |
| Wolke et al. (2017) | 0,033 | -0,005 | 0,070 | 1,711 | 0,087 |
| Wong et al. (2014) | 0,111 | 0,066 | 0,155 | 4,859 | 0,000 |
| | 0,058 | 0,026 | 0,090 | 3,545 | 0,000 |

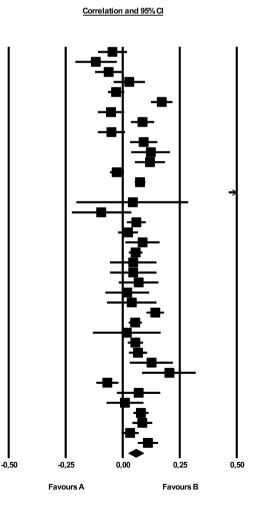


Fig. 2. Forest plot sex.

Meta Analysis

| Study name | | Statistics | s for each | study | | | Correlation and 95% Cl |
|--------------------------------|-------------|-----------------|----------------|----------------|----------------|-----|------------------------|
| | Correlation | Lower limit | Upper limit | Z-Value | p-Value | | |
| Alvarez et al. a (2015) | 0,043 | 0,026 | 0,061 | 4,817 | 0,000 | 1 1 | I I |
| Alvarez et al. b (2015) | 0,100 | 0,067 | 0,133 | 5,892 | 0,000 | | |
| Athanasiou et al.a (2018) | 0,024 | -0,064 | 0,111 | 0,533 | 0,594 | | |
| Athanasiou et al.b (2018) | -0,083 | -0,153 | -0,011 | -2,269 | 0,023 | | |
| Athanasiou et al.c (2018) | 0,050 | -0,026 | 0,126 | 1,296 | 0,195 | | |
| Athanasiou et al.d (2018) | 0,021 | -0,041 | 0,083 | 0,669 | 0,503 | | |
| Barboza (2015) | 0,840 | 0,833 | 0,848 | 91,369 | 0,000 | | Γ_ |
| Bauman et al.a (2013) | 0,072 | 0,053 | 0,091 | 7,429 | 0,000 | | |
| Bauman et al.b (2013) | 0,088 | 0,037 | 0,138 | 3,397 | 0,001 | | |
| Buelga et al. (2017) | 0,107 | 0,047 | 0,166 | 3,488 | 0,000 | | |
| Charalampous et al. (2018) | 0,054 | -0,012 | 0,121 | 1,602 | 0,109 | | ∔∎∓ |
| Chen et al. b (2018) | -0,026 | -0,056 | 0,004 | -1,710 | 0,087 | | |
| Choi et al. (2019) | 0,021 | -0.002 | 0,044 | 1,770 | 0,077 | | |
| DeSmet et al. (2014) | -0,119 | -0,195 | -0.041 | -2,999 | 0,003 | | |
| DeSmet et al. (2018) | -0.035 | -0.071 | 0.001 | -1,918 | 0,055 | 1 1 | |
| Erreygers et al. (2016) | 0.090 | 0.049 | 0,130 | 4,334 | 0,000 | | |
| Extremera et al. (2018) | 0,081 | 0,033 | 0,128 | 3,292 | 0,001 | | |
| Ferrer-Cascales et al. (2019) | 0,065 | 0,022 | 0,108 | 2,959 | 0,003 | | |
| Festl and Quandt (2016) | 0,061 | 0,015 | 0,107 | 2,613 | 0,009 | | |
| Gámez-Guadix et al. (2015) | 0,067 | -0,008 | 0,142 | 1,748 | 0,080 | | |
| Giménez et al. (2015) | 0,007 | -0.014 | 0,045 | 1,029 | 0,303 | | <u></u> _ |
| González-Cabrera a (2018) | 0,089 | 0.025 | 0,040 | 2,705 | 0,007 | | |
| González-Cabrera b (2018) | 0,067 | 0.002 | 0,131 | 2,017 | 0,044 | | |
| González-Cabrera et al. (2017) | 0,007 | 0,002 | 0,240 | 2,733 | 0,044 | | |
| Hempill and Heerde (2014) | 0,138 | -0,078 | 0,240 | 1.253 | 0,000 | | |
| Holfeld & Mishna a (2019) | 0,070 | -0,017 | 0,156 | 1,579 | 0,210 | | |
| Holfeld & Mishna b (2019) | 0,030 | -0.066 | 0,130 | 0,614 | 0,539 | | |
| Holfeld & Mishna c (2019) | 0,030 | -0,000 | 0,125 | 1,629 | 0,339 | | |
| Jang et al. (2014) | 0,090 | 0,018 | 0,130 | 20,190 | 0,103 | | |
| Jetelina et al. (2019) | 0,121 | 0,109 | 0,133 | 11,286 | 0,000 | | |
| Katz et al. (2019) | 0.007 | -0.142 | 0,207 | 0.090 | 0,000 | | |
| Lonigro et al. (2019) | 0,007 | | | | | | |
| Messias et al. (2014) | 0.038 | -0,013 0.021 | 0,133 0.054 | 1,620 4,418 | 0,105 0.000 | | |
| | -, | - , | -, | ., | -, | | |
| Olumide et al. (2016) | 0,128 | -0,001 | 0,254 | 1,942 | 0,052 | | |
| Quintana et al. (2018) | -0,064 | -0,112 | -0,016 | -2,597 | 0,009 | | |
| Shapka et al. a (2018) | 0,060 | -0,035 | 0,154 | 1,236 | 0,217 | | |
| Shapka et al. b (2018) | 0,120 | 0,040 | 0,199 | 2,931 | 0,003 | 1 1 | |
| Simckes et al. a (2017) | 0,339 | 0,002 | 0,607 | 1,972 | 0,049 | 1 1 | |
| Simckes et al. b (2017) | 0,561 | 0,478 | 0,634 | 10,918 | 0,000 | 1 1 | |
| Tsitsika et al. (2015) | -0,011 | -0,047 | 0,025 | -0,614 | 0,539 | 1 1 | · · · · · |
| Veiga et al. (2017) | 0,120 | 0,087 | 0,152 | 7,156 | 0,000 | 1 1 | |
| Waasdorp and Bradshaw (2015) | | 0,079 | 0,199 | 4,503 | 0,000 | 1 1 | |
| Wachs et al. (2018) | 0,084 | 0,041 | 0,127 | 3,802 | 0,000 | 1 1 | |
| Williford et al. a (2013) | 0,020 | 0,000 | 0,040 | 1,991 | 0,046 | 1 1 | |
| Williford et al. b (2013) | 0,020 | 0,000 | 0,040 | 1,991 | 0,046 | 1 1 | |
| Williford et al. c (2013) | 0,020 | 0,000 | 0,040 | 1,991 | 0,046 | 1 1 | |
| Williford et al. d (2013) | 0,010 | -0,010 | 0,030 | 0,996 | 0,319 | I I | |
| Wolke et al. (2017) | 0,059 | 0,022 | 0,096 | 3,091 | 0,002 | 1 1 | |
| | 0,094 | 0.031 | 0,157 | 2,910 | 0,004 | I I | |

Fig. 3. Forest plot age.

-0,50

-0,25

Favours A

PRISMA-Statement website,). It was decided to perform the Egger regression test. The results of the Egger regression in sex indicate the non-existence of bias: value for the intersection (B0) of 1.359993 with a 95% confidence interval (-1.98391; 4.70376), with t = 0.82262, gl = 39, the p value of 1 tail (recommended) is 0.20786 and the p value of 2 tails is 0.41573. For the age variable it obtains a value for the

intersection (B0) of -1.00089, with a 95% confidence interval (-2.96586, 0.96408), with t = 1.02724, gl = 43. The p-value of 1 tail (recommended) is 0.15503 and the p-value of 2 tails is 0.31005. The non-significance of both tails exposes the non-existence of publication bias. In addition, the standard error values of sex (1.65316) and age (0.97435) are close to the regression line, which reaffirms the

0,25

Favours B

0,50

0,00

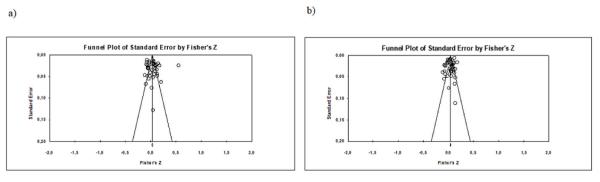


Fig. 4. Funnel plots for sex (a) and age (b).

non-existence of publication bias (Martin-Andrés & Luna del Castillo, 2004).

The funnel plots (Fig. 4) reaffirm the variability indicated by the Q and I statistics. However, their symmetry is consistent with the data provided by the Egger regression test. That is, regarding both the metaanalysis on sex (Fig. 4a) and age (Fig. 4b), the symmetry evidenced implies that both studies have captured all relevant studies. However, it should be noted that, in the case of sex (Fig. 4a), a study that stands out clearly on the right margin is Cuadrado-Gordillo and Fernández-Antelo (2016), which presents a moderate high correlation (0.502) between sex and rate of cyber-victimization in the Spanish sample. This aspect had already been evidenced in the graph of Forest Plot (Fig. 3).

3.4. Moderating variables

The evidence found on both variables could be modulated by other moderating factors such as culture (Botella & Sánchez, 2015), which is responsible for the high heterogeneity. In order to determine its role, a meta-regression test was carried out by applying a comparative model.

With regard to the meta-analysis on sex differences (see Table 3 and Fig. 4), the following models were generated. Model 1 (simple) and model 2 (continental culture), do not explain in percentage the variance. However, differences between cultures are observed (see Table 3) although no significant differences are established.

With regard to the meta-analysis on age differences (see Table 2 and Fig. 4), model 1 (simple) does not explain in any percentage the variability of the results. Model 2 (continental culture), explains 9%, but they are not statistically significant data (p > 0.05). However, in model 2 (see Table 3), as shown by the meta-regression, it is the American culture that is statistically significant compared to its counterparts (ST = -0.2134, z-value: -2.23 and p = 0.0257). In other words, American culture plays a key role in the age differences in the rate of cybervictimization (see Fig. 4). Consequently, the variability exposed by the Q, I2 statisticians and the Funnel Plot could be understood by this cultural diversity.

4. Discussion

The present study by means of meta-analysis and review of the literature investigates the incidence of sociodemographic variables such as sex and age in the rate of cyber-victimization. Similarly, it has been interested in the moderating role of culture according to the different continents. What is new is that, unlike previous studies which have focused their research on specific population nuclei or in any case on the comparison of different countries (Athanasiou et al., 2018), this one presents a variety of cultures grouped by continent.

In general, the results are consistent with previous research concluding that sex is a mild risk factor with little significance. Therefore, both sexes can be recipients of online attacks on equal terms, although there are slight indications of greater involvement of the female sex (Moreno-Ruiz et al., 2019). In this sense, if cyberbullying is considered a relational-indirect aggression, the female sex will be more likely to suffer from it (Barlett & Coyne, 2014).

However, the higher or lower prevalence of one sex in cybervictimization rates seems to be mediated by cultural factors. For example, in Australia the difference with women is not significant (Hemphill & Heerde, 2014). In North America, the general view, despite some exceptions (Holfeld & Mishna, 2019; Simckes, 2017), is that there are significant differences between the sexes, with women being more likely to be victims (Barboza, 2015; Bauman et al., 2013; Messias et al., 2014; Pettalia et al., 2013; Waasdorp & Bradshaw, 2015). The same is mostly true of European culture both in Nordic countries such as Finland (Williford et al., 2013) and in Mediterranean countries such as Spain and Italy (Álvarez-García et al., 2015; Baldry et al., 2017; Buelga et al., 2017; Cuadrado-Gordillo & Fernández-Antelo, 2016; Extremera et al., 2018; González-Cabrera et al., 2017, 2018; Moreno-Ruiz et al., 2019). However, there are also exceptions such as Portugal (Veiga Simão et al., 2017), Italy (Lonigro et al., 2015) or Israel (Katz et al., 2019). Eastern cultures are the only ones where homogeneity is found with a higher rate of cyber-victimization of males (Chen et al., 2018; Choi et al., 2019; Shapka et al., 2018; Wong et al., 2014) which is in line with earlier studies such as Chang et al. This peculiarity can be understood, for example, in Chinese society, where daughters are more vulnerable to family violence and this, in turn, represents one of the conditions for generating a future bully (People's Republic of China, 2005). Furthermore, concepts and meanings are applied differently in the East than in the West, as knowledge is historically elaborated and applied within very different socio-historical contexts (Kazue et al., 2006). On the other hand, it should be noted that privacy in the online environment is a determining element in cyber-victimization. In this sense, Festl and Quandt (2016) point to increased exposure to antisocial online activities and increased sexual risk in males.

On the other hand, age also has a small relationship with the rate of cyber-victimization with a declining trend that indicates that the younger the age, the greater the risk of cyberbullying (Lonigro et al., 2015; Shapka et al., 2018). Age emerges as a risk factor in cyberbullying related to early adolescence and associated with the greater impulsiveness of younger individuals, greater access to new technologies and more hours of internet use (Choi et al., 2019; Marciano et al., 2020). However, females are more likely to be cyberbullied in early adolescence, while males are more likely to be cyberbullied in late adolescence (Barlett & Coyne, 2014). Some research does not seem to find that cyber-victimization is age-mediated, as is the case with Katz et al. (2019) in Israel, or Buelga et al. (2017), Extremera et al. (2018), Ferrer-Cascales et al. (2019), Giménez Gualdo et al. (2015) in Spain. However, there are indications that age does play a determinant role in severe cases of cyber-victimization (Álvarez-García et al., 2015). The same diversity of results is presented by Asian, African (Olumide et al., 2015) or Australian cultures (Hemphill & Heerde, 2014). In the former, those who find a higher rate of cyber-victimization in early adolescence predominate (Chen et al., 2018; Jang et al., 2014) compared to those who point to an increase in the rate of cyber-bullying with age (Choi et al., 2019). In support of the results of this study, longitudinal research by Gámez-Guadix et al. (2015) agrees with authors such as Sakellariou et al. (2012) or Tokunaga (2010) on the existence of a curve whose peak is reached in middle adolescence, in the case of Spain between 15 and 16 years and in the United Kingdom between 13 and 14 years (Wolke et al., 2017).

The differences introduced by culture when assessing the results and the relevant role of the North American continental culture represented by Canada and the USA should also be highlighted. Different metaanalyses point out the significance between cyber-victimization and variables such as age or country (Barlett & Coyne, 2014; Chen et al., 2017). However, the present work, in addition to finding these direct

Table 2

| Model comparison according to random effects and Z-distribution mod | el. |
|---|-----|
|---|-----|

| | SEX | | | | | AGE | | | | |
|---|-------------------------|---------------------|--------------|--------|-------------------------|-------------------------|---------------------|---------------------|--------|---------------------|
| | TauS q | R ² | Q | gl | P-Value | TauS q | R ² | Q | gl | P-Value |
| Model 1 SIMPLE Model CULTURA CONTINENTAL | 0,0122 0,0134 | 0,00 0,00 | 0,00 1,78 | 0 4 | 1,0000 0,7770 | 0,0023 0,0021 | 0,00 0,09 | 0,00 9,34 | 0 5 | 1,00 0,09 |

Table 3

Meta-regression of sex and age by culture.

| Meta-regression SEX | | | | | | | | | |
|---------------------|-------------|----------------|-----------|------------|---------|-----------|------|----|------|
| Covariance | Coefficient | Standard error | 95% lower | 95% higher | Z value | 2-value p | Q | gl | Р |
| Intersection | -0,0862 | 0,1178 | -0,3171 | 0,1447 | -0,73 | 0,4643 | 1,78 | 4 | 0,77 |
| African Culture | 0,2137 | 0,1722 | -0,1239 | 0,5513 | 1,24 | 0,2147 | | | |
| Asian Culture | 0,1444 | 0,1295 | -0,1094 | 0,3981 | 1,12 | 0,2648 | | | |
| Culture European | 0,1228 | 0,1202 | -0,1128 | 0,3584 | 1,02 | 0,3068 | | | |
| Culture American | 0,1302 | 0,1255 | -0,1157 | 0,3762 | 1,04 | 0,2994 | | | |
| Meta-regression AGE | | | | | | | | | |
| Intersection | 0,1292 | 0,0809 | -0,0294 | 0,2878 | 1,60 | 0,1102 | 9,34 | 5 | 0,09 |
| African Culture | -0,1114 | 0,0826 | -0,2733 | 0,0505 | -1,35 | 0,1774 | | | |
| Asian Culture | -0,1256 | 0,0899 | -0,3018 | 0,0506 | -1,40 | 0,1625 | | | |
| Culture European | -0,0883 | 0,0816 | -0,2482 | 0,0717 | -1,08 | 0,2794 | | | |
| Culture American | -0,2134 | 0,0957 | -0,4009 | -0,0259 | -2,23 | 0,0257 | | | |
| Culture Oceanica | 0,0101 | 0,1450 | -0,2741 | 0,2942 | 0,07 | 0,9447 | | | |

relationships, concludes that culture acts as a moderator in the relationship between age and cyber-victimization. The theoretical review and the meta-regression show remarkable differences in the age variable attending to continental cultures in cyber-victimization. At the beginning, both point out that in the USA and Canada age plays a relevant role, highlighting that the younger the age, the higher the cyberbullying rate and the conformity of the data (Waasdorp & Bradshaw, 2015). In contrast, different studies show that the late teenage years are the ones with the highest rate of cyber-victimization, as it increases with age (Bauman et al., 2013; Messias et al., 2014; Shapka et al., 2018; Simckes et al., 2017). The European continent shows the intrinsic cultural diversity of each European country. In Belgium, similar data to the North American ones are obtained as more cases of cyber-victimization are observed in early adolescents (DeSmet et al., 2014, 2018; Erreygers et al., 2016). Therefore, and despite the small difference found on the basis of age, there is a great variety of results, being the North American culture the only one that contributes to explain the age differences in cyberbullying.

On the other hand, and despite the fact that they have not been the subject of statistical analysis, there are a series of variables that have been considered by the authors of the studies included in this metaanalysis such as family, individual differences and Internet consumption that deserve reflection and future study (Baldry et al., 2019; Baumgartner et al., 2014; Huang et al., 2020; Jetelina et al., 2019; Quintana-Orts & Rey, 2018; Turliuc et al., 2020).

5. Conclusions

The meta-analysis carried out has led to the conclusion that the effect sizes, although significant, are low for both sex and age. However, there is no unanimity in the results, as indicated by previous studies, on the role of women in cyber-victimization, although there are indications that women are more likely to suffer from cyber-bullying than younger individuals. The theoretical review and meta-regressions point to culture as a moderating element in explaining the rate of cybervictimization, albeit in low percentages. This would indicate the need to investigate, in the future, the role of other variables such as the type of use made of the internet, parental control, personality elements, etc., which would provide more conclusive data on the subject. This would either highlight the importance of these latter variables, or it would show the need to combine all of them to explain this problem.

Declaration of competing Interest

This manuscript has not been published and is not under consideration for publication elsewhere. We have no conflicts of interest to disclose.

Data availability

No data was used for the research described in the article.

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